

## The number of contributors

- Q: What's your opinion about how likely it is that there are more than four contributors to this mixed DNA sample
- A: I have absolutely no idea and nor does [the prosecution witness]. IN


## OUT



## Underestimating - $H_{p}$ true



One under

## Underestimating - $\mathrm{H}_{\mathrm{a}}$ true



## Overestimating $-H_{p}$ true

$4 p N$ vs $4 p N+1$
4p mixtures

Non-Conservative
O Series1
$\triangle$ Series2


One over

|  | True donors | False donors |
| :---: | :---: | :---: |
| One under | Kicks out the <br> smallest, you <br> didn't think it <br> was there | Some inc $\rightarrow$ excl |
|  |  |  |
|  |  |  |


|  | True donors | False donors |
| :---: | :---: | :---: |
| One under | Kicks out the <br> smallest, you <br> didn't think it <br> was there | Some inc $\rightarrow$ excl |
| One over | Big ones stay the <br> same smallest $\rightarrow$ <br> down 2-3 orders | Some excl $\rightarrow$ inc |
|  |  |  |


|  | True donors | False donors |
| :---: | :---: | :---: |
| One under | Kicks out the <br> smallest, you <br> didn't think it <br> was there | Some inc $\rightarrow$ excl |
| One over | Big ones stay the <br> same smallest $\rightarrow$ <br> down 2-3 orders | Some excl $\rightarrow$ inc |
| One over and $\mathbf{M x}$ <br> prior | Stay the same | Some excl $\rightarrow$ inc |

As long as your LR is big then you are correct or conservative

## Contents lists available at ScienceDirect

Forensic Science International: Genetics
journal homepage: www.elsevier.com/locate/fsigen

Research paper
Contributors are a nuisance (parameter) for DNA mixture evidence evaluation updates

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- "Thus...the LR will be determined as a weighted average of $L R(n)$ each with the same number $n$ in the numerator and in the denominator."


## Consider the weights

$$
w_{n}=\operatorname{Pr}\left(N=n \mid G_{P}, G_{C}, H_{a}\right) \frac{\operatorname{Pr}\left(N=n \mid H_{p}\right)}{\operatorname{Pr}\left(N=n \mid H_{a}\right)}
$$

where
$H_{\mathrm{p}}$ : The POI is a donor
$H_{\mathrm{a}}$ : The POI is not a donor
N is the number of contributors
$G_{\mathrm{C}}$ is the profile of the crime stain and
$G_{\mathrm{P}}$ is the profile of the person of interest.

## Consider the weights

$$
w_{n}=\operatorname{Pr}\left(N=n \mid G_{P}, G_{C}, H_{a}\right)
$$

You need to assign $N$
You know the POl's genotype
You know the crime sample
You assume POI is not a donor

- This suggests no justification
- To look at $G_{p}$ and $G_{c}$ and add 1 to "fit" $P$
- This suggests little justification for different N in numerator and denominator


## Verbal scales



## Adventitious matches can happen... and always could

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Forensic Science International: Genetics
journal homepage: www.elsevier.com/locate/fsigen

Research paper

# Internal validation of STRmix ${ }^{\mathrm{TM}}$ - A multi laboratory response to PCAST 

Jo-Anne Bright ${ }^{\text {a,* }}$, Rebecca Richards ${ }^{\text {a }}$, Maarten Kruijver ${ }^{\text {a }}$, Hannah Kelly ${ }^{\text {a }}$, Catherine McGovern ${ }^{\text {a }}$, Alan Magee ${ }^{\mathrm{b}}$, Andrew McWhorter ${ }^{\mathrm{c}}$, Anne Ciecko ${ }^{\mathrm{d}}$, Brian Peck ${ }^{\mathrm{e}}$, Chase Baumgartner ${ }^{\mathrm{f}}$, Christina Buettner ${ }^{g}$, Scott McWilliams ${ }^{g}$, Claire McKenna ${ }^{\text {h }}$, Colin Gallacher ${ }^{\text {i }}$, Ben Mallinder ${ }^{i}$, Darren Wright ${ }^{\mathrm{j}}$, Deven Johnson ${ }^{\mathrm{k}}$, Dorothy Catella ${ }^{1}$, Eugene Lien ${ }^{\mathrm{m}}$, Craig O'Connor ${ }^{\mathrm{m}}$, George Duncan ${ }^{\text {n }}$, Jason Bundy ${ }^{\text {o }}$, Jillian Echard ${ }^{\mathrm{p}}$, John Lowe ${ }^{\mathrm{q}}$, Joshua Stewart ${ }^{\mathrm{r}}$, Kathleen Corrado ${ }^{\text {s }}$, Sheila Gentile ${ }^{\text {s }}$, Marla Kaplan ${ }^{\mathrm{t}}$, Michelle Hassler ${ }^{\mathrm{u}}$, Naomi McDonald ${ }^{\mathrm{v}}$, Paul Hulme ${ }^{\mathrm{w}}$, Rachel H. Oefelein ${ }^{\mathrm{x}}$, Shawn Montpetit ${ }^{\mathrm{y}}$, Melissa Strong ${ }^{\mathrm{y}}$, Sarah Noël ${ }^{\mathrm{z}}$, Simon Malsom ${ }^{\mathrm{A}}$, Steven Myers ${ }^{\mathrm{B}}$, Susan Welti ${ }^{\mathrm{C}}$, Tamyra Moretti ${ }^{\mathrm{D}}$, Teresa McMahon ${ }^{\mathrm{E}}$, Thomas Grill ${ }^{\mathrm{F}}$, Tim Kalafut ${ }^{\mathrm{G}}$, MaryMargaret Greer-Ritzheimer ${ }^{\mathrm{H}}$, Vickie Beamer ${ }^{\mathrm{I}}$, Duncan A. Taylor ${ }^{\mathrm{J}, \mathrm{K}}$, John S. Buckleton ${ }^{\text {a,L }}$


## False inclusions (Adventitious



## Contents lists available at ScienceDirect

Forensic Science International: Genetics

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journal homepage: www.elsevier.com/locate/fsig
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Searching mixed DNA profiles directly against profile databases
Jo-Anne Bright ${ }^{\text {a,b,* }}$, Duncan Taylor ${ }^{\text {c }}$, James Curran ${ }^{\text {b }}$, John Buckleton ${ }^{\text {a }}$

Highest adventitious match 730,000

## False donor testing

- This tests known false donors against the profile
- Either use a database (say staff) or
- Simulate
- Run against the profile with your system,
- Record the results and present (?)
- Problem .... To test $L R=x$ you need at least $x$
- Turing informs us that an $L R$ of x happens less than 1 in $x$



## Internal validation compilation

2,825 mixtures $28,250,000$ false donors

| LR for $\mathrm{H}_{\mathrm{p}}$ Support and 1/LR for $H_{d}$ Support | Verbal Qualifier | Expected less than |
| :---: | :---: | :---: |
| [1-2) | Uninformative | 1 in 2 |
| [2-99) | Limited Support | 1 in 99 |
| [99-9999) | Moderate Support | 1 in 9,999 |
| [9999-999,999) | Strong Support | 1 in 999,999 |
| 2999,999 | Very Strong Support |  |

## Internal validation compilation

2,825 mixtures $28,250,000$ false donors

| LR for $\mathrm{H}_{\text {p }}$ |  |  |  |
| :---: | :---: | :---: | :---: |
| Support and |  | Expected | Fraction of false donor LRs |
| $\text { 1/LR for } \mathrm{H}_{d}$ | Verbal Qualifier | less than | in this range ( $\mathrm{N}=$ |
| Support |  |  | 28,250,000) |
| [1-2) | Uninformative | 1 in 7 | 1 in 312 |
| [2-99) | Limited Support | $1 \text { ing9 }$ | 1 in 318 |
| [99-9999) | Moderate Support | 1 in 9,999 | 1 in 18,000 |
| [9999-999,999) | Strong Support | 1 in 999,999 | 1 in 1,400,000 |
| <999,999 | Very Strong Support |  | $\bigcirc$ |



# Juror comprehension of forensic expert testimony: A literature review and gap analysis 

## Heidi Eldridge

RTI International 3040 E Cornwallis Rd, Research Triangle Park, NC, 27709, USA

From Turing we can infer that

$$
\tilde{p} \leq \frac{1}{L R_{P O I}}
$$

Equation 2

The chance of an $L R$ greater than or equal to $L R_{\text {POI }}$ is less than $1 / L R_{\text {POI }}$ This is true for every $L R$ not just $L R_{\text {Pol }}$

## The distribution of Ha true

-the shape depends on the profile -there will be a maximum,

- Not directly known to us but potentially calculable
- this is probably slightly bigger that the largest Hp true

A low level four person mixture ( $4: 3: 2: 1 \mathrm{pg}$ ) 12 loci where


Contents lists available at ScienceDirect

## Forensic Science International: Genetics

journal homepage: www.elsevier.com/locate/fsig

Research paper
Importance sampling allows $H_{\mathrm{d}}$ true tests of highly discriminating DNA profiles

Duncan Taylor ${ }^{\text {a,b,* }}$, James M. Curran ${ }^{\text {c }}$, John Buckleton ${ }^{\text {d,e }}$

| 400000 | Glok Propositions |
| :---: | :---: |
|  |  |
| 0 | Prosecution: Simulated profile + Unknown |
| 350000 | Defence: Two unknowns |
| 250000 | Actual simulations Equivalent 'naïve' |
| 150000 | 1000 1.12 $\times 10^{21}$ |
| 100000 | Average $H_{a}$ true LR |
| 50000 | Hher 1.12 |
|  |  <br>  |

$\log _{10}(L R)$



## Contents lists available at SciVerse ScienceDirect

## Forensic Science International

# Reliable support: Measuring calibration of likelihood ratios ${ }^{\text {T }}$ 

## Daniel Ramos*, Joaquin Gonzalez-Rodriguez

Research Institute on Forensic Science (ICFS), ATVS, Biometric Recognition Group, Escuela Politecnica Superior, Universidad Autonoma de Madrid, C/ Francisco Tomas y Valiente 11, E-28049 Madrid, Spain

## Calibration


$50 \%$ chance of rain
How could we check
Collect times he says $50 \%$ and see if $50 \%$ of these have rain


## In the 31 lab set

- 28,250,000 false donors
- 10,297 true donors
- Prior odds 10,297/28,250,000
- Each LR can be turned into a posterior odds and
- Then into a posterior probability
- Are we "right" the "right" number of times

| Range of posterior probabilities | Observed | Number of false donors in this interval | Number of true donors in this interval |
| :---: | :---: | :---: | :---: |
| 0.977 to 1.000 |  | 4 | 7657 |
| 0.891 to 0.977 |  | 7,657 |  |
| 0.813 to 0.891 |  | $\overline{4+7,657}=0.9995$ |  |
| 0.398 to 0.813 |  |  |  |
| 0.158 to 0.398 |  | 200 | 201 |
| 0.063 to 0.158 |  | 563 | 162 |
| 0.025 to 0.063 |  | 1,366 | 165 |
| 0.010 to 0.025 |  | 3,533 | 131 |
| 0.004 to 0.010 |  | 9,569 | 133 |
| 0.002 to 0.004 |  | 24,603 | 115 |
| 0.001 to 0.002 |  | 64,106 | 112 |
| 0.0003 to 0.0006 |  | 156,994 | 115 |
| 0.0000 to 0.0003 |  | 28,037,070 | 760 |


| Range of posterior probabilities | Observed | Number of false donors in this interval | Number of true donors in this interval |
| :---: | :---: | :---: | :---: |
| 0.977 to 1.000 | 0.9995 | 4 | 7657 |
| 0.891 to 0.977 |  | 5 | 342 |
| 0.813 to 0.891 |  | 7 | 111 |
| 0.398 to 0.813 |  | 60 | 314 |
| 0.158 to 0.398 |  | 200 | 201 |
| 0.063 to 0.158 |  | 563 | 162 |
| 0.025 to 0.063 |  | 1,366 | 165 |
| 0.010 to 0.025 |  | 3,533 | 131 |
| 0.004 to 0.010 |  | 9,569 | 133 |
| 0.002 to 0.004 |  | 24,603 | 115 |
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| Range of posterior probabilities | Observed | Number of false donors in this interval | Number of true donors in this interval |
| :---: | :---: | :---: | :---: |
| 0.977 to 1.000 | 0.9995 | 4 | 7657 |
| 0.891 to 0.977 | 0.9856 | We are "right" too often Not enough false donors up here |  |
| 0.813 to 0.891 |  |  |  |
| 0.398 to 0.813 |  |  |  |
| 0.158 to 0.398 |  | 200 | 201 |
| 0.063 to 0.158 |  | 563 | 162 |
| 0.025 to 0.063 |  | 1,366 | 165 |
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| 0.0000 to 0.0003 |  | 28,037,070 | 760 |


| Range of <br> posterior <br> probabilities | Number of <br> false donors in | Number of true <br> donors in this <br> interval |  |
| :--- | ---: | ---: | ---: |
| $\mathbf{0 . 9 7 7}$ to 1.000 | 0.9995 | 4 | 7657 |
| $\mathbf{0 . 8 9 1}$ to 0.977 | $\mathbf{0 . 9 8 5 6}$ | 5 | 342 |
| $\mathbf{0 . 8 1 3}$ to 0.891 | $\mathbf{0 . 9 4 0 7}$ | $\mathbf{7}$ | 111 |
| $\mathbf{0 . 3 9 8}$ to 0.813 | $\mathbf{0 . 8 3 9 6}$ | 60 | 314 |
| $\mathbf{0 . 1 5 8}$ to 0.398 | $\mathbf{0 . 5 0 1 2}$ | 200 | 201 |
| $\mathbf{0 . 0 6 3}$ to 0.158 | $\mathbf{0 . 2 2 3 4}$ | 563 | 162 |
| $\mathbf{0 . 0 2 5}$ to 0.063 | $\mathbf{0 . 1 0 7 8}$ | 1,366 | 165 |
| $\mathbf{0 . 0 1 0}$ to 0.025 | $\mathbf{0 . 0 3 5 8}$ | 3,533 | 131 |
| $\mathbf{0 . 0 0 4}$ to 0.010 | $\mathbf{0 . 0 1 3 7}$ | 9,569 | 133 |
| $\mathbf{0 . 0 0 2}$ to 0.004 | $\mathbf{0 . 0 0 4 7}$ | 24,603 | 115 |
| $\mathbf{0 . 0 0 1}$ to 0.002 | 0.0017 | 64,106 | 112 |
| $\mathbf{0 . 0 0 0 3}$ to 0.0006 | $\mathbf{0 . 0 0 0 7}$ | 156,994 | 115 |
| $\mathbf{0 . 0 0 0 0}$ to 0.0003 | 0.000027 | $28,037,070$ | 760 |

## Communication



Report of Professor Allan Jamieson in the case of Donte Lee
$8^{\text {th }}$ May 2017
Occupation: Director of The Forensic Institute
No one understand the LR

This illustrates that if the LRs of all the millions of potential genotypes from a mixture were calculated and then arranged in order of size, the suspect is unlikely to be the highest LR.
In other words, the LR provides only the weight of evidence against the specific defendant without reference to other people who would also have a LR greater that 1 (i.e. support for the prosecution hypothesis).
In effect, the LR is a sophisticated version of the disparaged 'consistent with' statement.

All profiles





## Weights and ranks



Number of genotypes STRmix considered

Number of genotypes at this locus


In our example $8.55 \times 10^{38}$ genotypes
$7.5 \times 10^{9}$ people
Only about 1 in $10^{29}$ genotypes exist
There are about $6 \times 10^{7}$ genotypes above our rank
Hence potentially no actual people above our rank

Most genotypes do not exist

Weir, BS

## Likelihood ratio

"The probability of observing this evidence is $n$ times more likely if it arose from $\mathrm{Mr} \mathrm{X}+\mathrm{an}$ unknown person rather than two unknowns"


- Is NOT measuring the probability of Mr Lee being a contributor - many profiles will produce a high LR
- High LRs can be obtained for false propositions
- Depends on the number of contributors
- Does not test all of the possible explanations for the evidence

A statement about the probability that Mr Smith left the stain can only be made from all the evidence, not from the DNA alone.

The DNA evidence by itself increases the odds that Mr Smith is the donor LR times
Over what they would be from the other evidence

This represents extremely strong support that he is the donor DNA evidence

## Its not quite a $p$-value. But there has been considerable criticism.

The LR is the best summary of the evidence.

## Rewording

- Concern about pseudo-frequentist expression
- Other profiles "near-by"
- Loss of posterior = LR x prior construct
- Trying to "not change"


## Theory of communication - Q\&A

- No feedback with jury
- Deliberate noise

Field o


## Saying it better does not guarantee understanding better

Bernard Robertson I G. A. Vignaux | Charles E. H. Berger
Interpreting
Evidence

Second Edition


They have some great analogies
N But they are long and I suspect you cannot do them in court

## Feedback Receiver

## Saying it better does not guarantee understanding better

## Field of Exp

## I REALLY MEAN

Sender
Decoder

Feedback Receiver

## Conclusion

- Number of Contributors
- LR stable over NoC
- Adventitious matched do happen - Actually at less than the expected rate
- Reliable support - calibration
- Communication - cannot be fixed at one end


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- Professor Bruce Budowle
- Lynn Garcia

